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*Wages, employment and
economic shocks:
Evidence from Indonesia*

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Abstract

After over a quarter century of sustained economic growth, Indonesia was struck by a large and unanticipated crisis at the end of the 20th Century. Real GDP declined by about 12% in 1998. Using 13 years of annual labor force data in conjunction with two waves of a household panel, the Indonesia Family Life Survey (IFLS), this paper examines the impact of the crisis on labor market outcomes. Aggregate employment has remained remarkably robust through the crisis although there has been significant switching within sectors. The drama of the crisis lies not in aggregate employment but in real hourly earnings which, in one year, collapsed by around 40% for urban workers -- be they males or females, in the market sector or self-employed. Declines of the same magnitude are recorded for females in the rural sector and rural males working for a wage. In stark contrast, real hourly earnings of self-employed males in rural areas have remained essentially stable. Given these workers account for about one quarter of the male work force in Indonesia, conclusions about the effects of the crisis that focus only on the market wage sector substantially overstate the magnitude of the crisis. Indonesian families have displayed a remarkable capacity for resilience in the face of the crisis. We estimate that declines in real household incomes are about half the magnitude of the declines in individual hourly earnings indicating that households have adopted an array of strategies to mitigate the effects of the crisis.

Introduction

After over a quarter century of sustained economic growth, Indonesia was struck by a major economic crisis at the end of the 20th Century. The crisis, which worked its way through many of the South East Asian "tigers" was presaged by the collapse of the Thai baht in the middle of 1997. During the second half of the year, the Indonesian central bank attempted to stave off pressure on its currency, the rupiah. Interest rates quadrupled and the rupiah depreciated by about 60%. In January 1998, the rupiah collapsed. It lost about 75% of its value in a matter of a few days. The collapse of the rupiah was followed by spiralling prices; annual inflation is estimated to be about 80% for 1998.

The impact of the crisis in Indonesia was not confined to the financial sector. The economy went into a tailspin. Real output in 1998 was estimated to have been about 12% below its 1997 level (a bad year for some Indonesians because of the drought due to El Nino) and economic growth turned positive only in 2000. The crisis was accompanied by dramatic and far-reaching political change.

By any metric, the crisis in Indonesia has been extremely severe -- far more severe than in any of its South East Asian neighbors. For example, a year after the crisis began, the Thai baht and Philippine peso stood at about half their pre-crisis values; the Indonesian rupiah had fallen to about one-quarter its 1997 value. A collapse of this magnitude is not unique: the Russian rouble and Ecuadoran sucre were both reduced to one-quarter of their value over twelve months in 1998/1999. Relative to the Russian and Ecuadoran crises, the crisis in Indonesia was different in one key dimension: its timing and virulence were largely unanticipated. For example, in January 1998, the Government of Indonesia predicted that for the year economic growth would be zero and inflation would be 20%. They were not alone in painting this rosy picture. In July, 1998, the World Bank admitted, "Few predicted the Indonesian crisis, and certainly none its severity.... Indeed, Indonesia seemed better positioned than many other crisis countries on bellwether indicators such as growth, the fiscal stance, the external current account, foreign exchange reserves, and inflation." (World Bank, 1998.)

The basic facts of the Indonesian crisis have been widely documented at a macro level (Cameron, 1999, Pangestu, 1999). The underlying causes of the crisis, and appropriate macroeconomic remedies, have been the subject of considerable debate. (See, for example, Corsetti, Pesenti and Roubini, 1998a, 1998b; Radelet and Sachs, 1998; Krugman, 1998; and IMF, 1999.)

Substantially less is known about the magnitude and distribution of the micro-economic impacts on the lives of Indonesians. For example, there is controversy about even basic questions regarding the magnitude of the shock on wages and employment. Some have claimed that the crisis has brought rampant unemployment (ILO, 1999) and called for investments in public works programs (World Bank 1998). Others have suggested that employment remained remarkably stable through the crisis while real wages collapsed (Frankenberg, Thomas and Beegle, 1999; Papanek and Handoko, 1999, and Feridhanusetyawan, 1999).

Many key issues regarding the effects of economic shocks on labor markets have received little attention in the literature. How do labor markets adjust to a large, abrupt economic decline of the magnitude experienced in Indonesia? Are urban and rural labor markets affected similarly? Do the economic effects differ among the lower, middle, and higher income segments of the labor market? Answering these questions is important not only for documenting the size of the Indonesian crisis and its distributional impact, but, also to provide insights into more general questions about the functioning of markets -- particularly labor markets -- in low income settings.

While the questions are important, answers are in short supply especially in developing countries where the social science data infrastructure typically has not been in place to monitor the labor market effects of large economic shocks.¹ Fortunately, in this dimension Indonesia represents an unusual counter-example. First, since the mid-1980s, Indonesia has conducted excellent large-scale, cross-section, nationally representative employment and income surveys, *Survei Angkatan Kerja Nasional* (SAKERNAS); those surveys are currently conducted annually. Second, during the years proceeding and in the midst of the crisis, a longitudinal survey, the Indonesian Family Life Survey (IFLS) was fielded. The IFLS is an ongoing survey of individuals, families, and communities in Indonesia that offers unusually rich detail about many aspects of the lives of the respondents including their labor market activities. The first wave, IFLS1, was fielded in the second half of 1993, well before the crisis. The second wave, IFLS2, was conducted four years later, in 1997, as the currency crisis was beginning to unfold. Fieldwork for IFLS2 was completed prior to the collapse of the economy and before prices spiralled. In an effort to respond to the scientific and policy needs for quality and timely information on the impact of the crisis, a third round, IFLS2+, was fielded in late 1998, a year after IFLS2. IFLS2+ re-interviewed respondents from a 25% sub-sample of the original enumeration areas. This paper draws on the sample of almost 2,000 IFLS households who were interviewed both in late 1997 and about a year later in 1998. The IFLS data are used in conjunction with 13 years of the SAKERNAS to place the crisis in temporal context and document the nature of the changes that have taken place in Indonesian labor markets in recent years. By laying out these facts, we also lay the groundwork for implementing tests of hypotheses regarding the efficiency and completeness of markets in low income populations.

The remainder of the paper is divided into five sections. Section 1 describes some salient dimensions of the social, demographic and economic changes that have been taking place in Indonesia both over the long-term and during the economic decline of the late 1990s. The second section summarizes the main features of the data used in this research. Section 3 focusses on labor markets and describes the principal long term employment and wage trends in urban and rural labor markets as well as how those

¹See Fallon and Lucas (1999) for an excellent discussion of the very limited evidence on labor market effects of recent economic crises.

trends were disrupted by the economic crisis. The next section broadens the analysis to include incomes of the self-employed and highlights the role of dynamics in the labor market; we also examine the role of the family in adjusting to the crisis. The final section highlights the main conclusions.

Section 1. The Indonesian Context

Indonesia, the fourth most populous nation in the world, is an archipelago whose 13,000 islands are home to many different ethnic groups. The country varies a great deal in its urban and rural settlements. For example, population densities range from five people per square kilometer in the province of Irian Jaya to more than 700 people per square kilometer in Yogyakarta. The capital city, Jakarta, has been becoming increasingly important over time as the economic and political center of the country. Jakarta has been a major destination for permanent and circulatory migrants in the last three decades accounting for a quarter of interprovincial in-migrants even though it represents less than 5% of the total population. While not on the same scale as Jakarta, urban centers on other islands also grew in terms of their economic activity and the size of their populations. Still, in spite of the growth of a number of urban centers, Indonesia remains a largely agricultural country. In 1986, a fifth of Indonesians worked in urban areas; this fraction had risen to a third by 1997.

Thirty years ago, Indonesia was one of the poorest countries in the world. Until the recent financial crisis, it enjoyed high economic growth rates and was on the verge of joining the middle income countries. Prior to the crisis, in terms of size, its economy was comparable to that of Malaysia, the Philippines and Thailand *combined*. On average, GNP per capita in Indonesia grew by 4.5% per annum from the mid-sixties until 1998. However, economic growth was far from uniform across the country with economic heterogeneity having, if anything, increased over time. With the growth of industry and the service sector, the relative importance of agriculture gradually declined (from 55% of total employment in 1986 to 41% by 1997). Over the same years, industrial employment more than doubled from 8% to 19% of the labor force. Not surprisingly, employment in the formal wage sector was expanding, rising from a quarter to a third of all jobs during the same years. This expansion of the formal wage sector, notwithstanding, it is important to emphasize that at the start of the economic crisis, most Indonesians workers were either self-employed or unpaid family workers.

Not only has the economy grown rapidly over the last quarter century but there has also been dramatic demographic and social change. Fertility rates have declined substantially and there have been massive investments in human capital. This is reflected, for example, in secondary school enrollment rates which have risen from a mere 6% in 1960 to more than 50% today; primary school enrollment is essentially universal. The investments are also reflected in the health of the population: at birth, the average Indonesian expects to live to age 61 today which is 50% longer than he or she expected 30 years

ago. Thus, the labor force has not only grown dramatically in size but also in quality. The consequences of these improvements are reflected in significant rising real wage levels over the last three decades.

Pride in its past economic achievements and optimism about its future were suddenly challenged by the economic crisis which was accompanied by dramatic shifts in the economic and political landscape of the country. As indicated in Figure 1, the rupiah came under pressure in the last half of 1997 when the exchange rate began showing signs of weakness. After falling by half from around 2,400 per US\$ to about 4,800 per US\$ by December 1997, the rupiah collapsed in January 1998 when, over the course of just a few days, the exchange rate fell by a factor of four to Rp16,000 per US\$. Although it soon recovered, by the middle of 1998, the rupiah had slumped back to the lows of January 1998. Since then, the rupiah has continued to oscillate, albeit at a lower amplitude and frequency. The extremely volatile exchange rate has contributed to considerable uncertainty in financial markets. This is reflected in interest rates which quadrupled in August 1997 and were subsequently very volatile. The banking sector fell into disarray and several major banks have been taken over by the Indonesian Bank Restructuring Agency. Turmoil in the financial sector has created havoc with both the confidence of investors and with the availability of credit.

Prices of many commodities spiraled upwards during the first three quarters of 1998. Annual inflation is estimated by the *Badan Pusat Statistik* (BPS), the central statistical bureau, to be about 80% for 1998. In part, this reflects the fact that subsidies were removed on several goods — most notably rice, oil and some fuels. Part of the inflation can be attributed to the fact that rice accounts for a substantial fraction of the average Indonesian's budget and its price more than doubled. Since the share spent on rice is greatest for the poorest, inflation likely had a bigger impact on the purchasing power of the poorest. Off-setting that effect, however, is the fact that some of the poorest are rice producers and as the price of food rose, so (net) food producers have benefitted from the improvement in their terms of trade. More generally, there is abundant evidence that price increases have not been uniform -- either across goods or across space. Inflation estimates produced by BPS are based on prices collected in urban areas; there are reasons to expect prices in rural areas have not moved in lock step with urban prices. Given the magnitudes of the swings in relative prices, it is likely that households will have responded to the price changes by substituting away from goods that have become relatively expensive towards goods whose prices have increased less. These behaviors are not reflected in the BPS inflation estimates (which are based on a basket of commodities calculated in 1996) although evidence from IFLS indicates that, in fact, households throughout the income distribution have altered their spending patterns in response to the relative price changes. Since the BPS province-specific price series is the only series that spans the time period considered in this paper, for the sake of consistency, incomes are deflated with that series throughout. The issues raised above should be kept in mind when interpreting the evidence. (For further discussion, see Thomas, Frankenberg, Beegle and Teruel, 1999, and Berry and Levinsohn, 1999.)

The hypothesis that net food producers may have been partially protected from the effects of the crisis needs to be tempered since a severe drought immediately preceded the financial crisis and it affected agriculture in many parts of the country - particularly in the east. Country-wide, rice production fell by 4% in 1997 with rice and soybeans being imported. Moreover, unusually severe forest fires raged in parts of Sumatra, Kalimantan, and Sulawesi affecting many aspects of economic life including agriculture and tourism.

Indonesia has also undergone dramatic transformation in the political sector. After thirty years of power, Suharto resigned from the Presidency in May 1998 in the midst of violent demonstrations throughout the country. Multi-party elections in 1999, diminished the influence of the military in political affairs and the granting of independence to East Timor may herald the re-birth of democracy in Indonesia although several other parts of the country are agitating for autonomy and political, ethnic and religious tensions continue to rock the nation.

Few Indonesians have been untouched by the crisis. For some, the impacts have been devastating, but for others, the crisis has likely brought new opportunities. Exporters, export producers and food producers may have fared far better than those engaged in the production of services and non-tradeables or those on fixed incomes. Moreover, some individuals may have more quickly embraced new opportunities or more efficiently offset the effects of the economic shocks that they have faced. Given the complexity and multi-faceted nature of the crisis, it is only with sound micro-level empirical evidence that it becomes possible to fully explore the nature and extent of behavioral responses by individuals and families to the crisis and, thereby, characterize with much confidence what the combined impacts of the various facets of the crisis have been and how they have varied across socioeconomic and geographic strata. Moreover, the massive upheavals in the Indonesian economy -- and the diversity of their impact -- provides an unparalleled opportunity to better understand the dynamics of urban and rural labor markets in low income settings as well as mechanisms used by families to smooth out the effects of large, unanticipated shocks.

Section 2. Data Sources

This study relies on survey data from two complementary sources: SAKERNAS and IFLS. IFLS is a large-scale integrated socioeconomic longitudinal survey containing extensive information on the lives of respondents, their households and families, as well as on the communities in which they live. The first wave was conducted in 1993 (IFLS1), with a follow-up in 1997 (IFLS2) and a special follow-up (of a 25% sub-sample) in late 1998 (IFLS2+).

IFLS

The IFLS1 sample is representative of about 83% of the Indonesian population living in 13 of the 27 provinces in the country.² Of the households included in the original listing for IFLS1, more than 93% were interviewed yielding a sample of 7,224 households. IFLS1 was fielded between August 1993 and January 1994. In each household, representative members (typically the female and male household heads) provided detailed household-level demographic and economic information. In addition to these respondents, several household members were also randomly selected and they were each asked about their own lives. These "main" respondents provided detailed information on a broad array of topics including education; migration; assets and wealth; use of health care and health status; marriage; fertility and contraception. The work history module contains questions on many dimensions of the respondents' current and prior labor market behavior, including income from earnings, hours of work and type of work in both their primary and secondary jobs. Questions are asked about wage employment, self-employed activities and work in family businesses. The work histories span the 5 years prior to the survey.

The second wave of the IFLS (IFLS2) was fielded between August 1997 and February 1998. Most of the interviews were completed by December 1997; the first two months of 1998 were spent tracking down movers who had not already been found. (Frankenberg and Thomas, 2000, describe the survey.) Our goal was to recontact all the original IFLS households and re-interview all IFLS1 "main" respondents (as well as all respondents born prior to 1967 independent of whether or not they were "main" respondents); we will refer to these as our "target" respondents. If a "target" respondent had moved, information was sought about the person's new location and, if that location was in any of the 13 IFLS provinces, attempts were made to interview the respondent in the new location. Excluding the households in which everyone had died (mostly single-person households), 94% of the IFLS1 households were re-interviewed and individual interviews were completed with 93% of target respondents. In IFLS2, household-level information was usually provided by the household head and spouse and we attempted to complete an individual interview with every household member; over 33,000 individuals were interviewed (representing over 95% of all household members). There are 7,600 households in IFLS2, the increase relative to IFLS1 arises because respondents who had split off from the original household and set up their own households were followed.

IFLS2+ was fielded a year after IFLS2, between August and December, 1998. A key goal was to provide insights about the likely immediate effects of Indonesia's economic crisis by collecting high quality, timely data on who had been affected and on the strategies individuals, households and communities adopted to mitigate the impact. Because there was neither the time nor resources to mount a survey of the

²The sample includes four provinces on Sumatra (North Sumatra, West Sumatra, South Sumatra, and Lampung), all five of the Javanese provinces (DKI Jakarta, West Java, Central Java, DI Yogyakarta, and East Java), and four provinces covering the remaining major island groups (Bali, West Nusa Tenggara, South Kalimantan, and South Sulawesi). The IFLS1 sampling scheme balanced the costs of surveying the more remote and sparsely-populated regions of Indonesia against the benefits of capturing the ethnic and socioeconomic diversity of the country.

same magnitude as IFLS2 (which took more than two years to plan and test), a scaled down survey was administered, while retaining the intellectual core of IFLS2. In IFLS2+, a 25% subsample of the IFLS enumeration areas (EAs) were selected and all households who had lived in those EAs in 1993, including split-offs followed in 1997, were included in the target sample. The EAs were chosen to be representative of the entire IFLS sample.³

Counting all original households in IFLS1 and the split-offs in IFLS2, there are 2,066 households in the IFLS2+ target sample. In spite of the turmoil in Indonesia during 1998, the fieldwork was remarkably successful. We relocated and re-interviewed more than 95% of the target households (including 60% of the IFLS1 households that had not been interviewed in IFLS2). The focus in this paper is on a comparison of behaviors in IFLS2, with behaviors in IFLS2+, about a year later. Taking as our sampling frame those households interviewed in 1997, over 98.5% were re-interviewed in 1998 and 95% of the 1997 respondents completed individual interviews in 1998. These "panel" respondents are the core of the data for our analyses based on the IFLS.⁴ Because the same respondents were interviewed twice, once just prior to the collapse of the rupiah in January 1998, and once during the crisis, a comparison of IFLS2 and IFLS2+ allows us to examine the dynamics of the labor market behaviors of individuals and families as the crisis progressed. (See Thomas, Smith and Frankenberg, 2000, for a detailed discussion of attrition.)

SAKERNAS

SAKERNAS is the principal Indonesian Labor Force survey administered by the BPS to compile and monitor provincial employment statistics. It has been conducted since 1986. Prior to 1995, the survey was conducted quarterly and annually in August after 1995. The survey covers a nationally representative sample of approximately 40,000 working-age individuals. The survey has been conducted every year except 1995 when the *Survei Penduduk Antar Sensus* (Intercensal demographic survey, SUPAS) was fielded instead. SUPAS covered the same questions on labor force activities as SAKERNAS on a sample that was three times bigger than SAKERNAS and so it is included in our time series of cross-section surveys.

The surveys record demographic information, including age, sex and marital status, of all household members. For those 10 years of age and older, information is collected on educational attainment and activities during the previous week. For those who were working, questions are asked about days and

³The sample was drawn in two stages. To reduce costs, 7 of the 13 IFLS provinces were selected: 2 on Sumatra (North and South Sumatra), 3 on Java (DKI Jakarta, West and Central Java), West Nusa Tenggara and South Kalimantan. These provinces span the full spectrum of socio-economic status and economic activity in the fuller IFLS sample. Second, within those provinces, we purposively drew 80 EAs with weighted probabilities to match the IFLS sample as closely as possible.

⁴IFLS2 contains over 21,000 individuals who completed the work history module. Of them, 5,345 respondents are included in the IFLS2+ target sample based on their EA of residence in 1993. 7,500 respondents completed the work history module in IFLS2+; of them, 2,434 are new respondents (not interviewed in 1997) and 5,067 are panel respondents interviewed in both 1997 and 1998.

hours of work as well as occupation and industry of employment. Monthly and weekly earnings are collected for all 'employees' or wage earners in the household.

SAKERNAS has several key advantages for our purposes. First, it provides an excellent base to lay out the main facts about labor market changes in Indonesia in the last 15 years. It will be used to cross-validate the main patterns of labor market changes observed in IFLS2/2+. Second, SAKERNAS spans a long enough time period so that the 1998 economic crisis can be given some temporal and economic perspective. Third, sample sizes are large so that small changes in outcomes can be measured precisely. The two principal disadvantages of SAKERNAS are that it contains no longitudinal information so that change at the individual or household level cannot be measured directly and it collects no information on incomes earned from self-employment. By drawing on IFLS2/2+, these disadvantages will be addressed.

A potentially important difference between the IFLS and SAKERNAS involves the extent of their geographical representation. While SAKERNAS is nationally representative, IFLS1 is limited to 13 provinces and IFLS2+ was further restricted to 7 of those provinces. These restrictions are far less severe than a simple province count would indicate since the most populous provinces were all part of IFLS1, IFLS2, and IFLS2+. Nonetheless, a legitimate question does arise about how comparable the two sets of surveys will be in their depiction on the economic crisis in light of their different geographical coverage.

To address this question, Table 1 lists some key labor market indicators on which we will rely heavily in the rest of the paper. The table uses SAKERNAS to measure the rate of growth of employment (by sector of work) and hourly wages in the formal (or wage employment) sector between 1986 and 1997 (in the left hand panel) and between 1997 and 1998 (in the right hand panel). The first column in each panel includes all 27 provinces in Indonesia;⁵ the second column restricts the SAKERNAS data to the 13 IFLS provinces and the third column to the 7 IFLS2+ provinces.

Whether changes in wages or employment rates are considered, the general characterization of growth in Indonesia in the decade preceding the crisis or of the labor market consequences of the crisis are virtually identical no matter whether the complete or more limited set of IFLS2+ provinces are used. While reassuring, this result should not be surprising since over 80% of the Indonesian population is represented in the IFLS sample frame and the additional geographic restrictions necessary for IFLS2+ were explicitly chosen to maintain the representativeness of the sample.

Consider first real wage changes in the formal wage sector.⁶ Between 1986 and 1997, mean real wages of males expanded by slightly more than 40%. This increase over more than a decade was almost

⁵The estimates include East Timor which was part of Indonesia at the time.

⁶Recall self-employment income is not recorded in SAKERNAS. Wages are computed as the sum of monthly cash and in-kind income earned divided by 4.33 times weekly hours. All wages are in terms of 1997 rupiah deflated using province-specific consumer price indices published by BPS.

entirely wiped out in one year during the economic crisis. Long-term wage growth of males was only slightly lower in the urban sector, but the urban sector was hit more severely by the economic shock. Between 1997 and 1998, wages of males declined by about 5% more in urban areas so that real wages were actually lower in 1998 than in 1986. In the formal wage sector, women had actually enjoyed greater wage increases in the years before the crisis. Between 1986 and 1997, wages of females rose by more than 60% thereby reducing the female/male wage gap in paid employment by more than 20%. The crisis did not discriminate across gender lines as female wages fell by about the same percentage as male wages. As was true for men, female wages declined more sharply in the urban areas -- about 8% more than in rural areas. During the growth years between 1986 and 1997, female wage growth had been 10 to 15% higher in urban wage jobs than in the rural wage jobs.

In contrast with wages, the overall fraction of the population in the labor force changed very little during the decade prior to the crisis and during the crisis. What action there is lies in movement between sectors of employment. During the growth years, male and female employment rates were increasing in the wage sector and declining in unpaid family jobs. The 1986-1997 growth in wage employment was about 3 percentage points higher for men than it was for women while the decline in unpaid family jobs was about two percentage points greater for women compared with men (of whom very few were unpaid family workers). The principal gender difference was that men were moving out of self-employed jobs while this sector remained relatively stable among women. These broad trends suggest that the main employment impact of the crisis was to shift male employment out of the wage sector into self-employment which, we will see below, is a shift that had already begun before the crisis.

For this paper, the key point emerging from Table 1 is that the employment trends and changes in wage rates are virtually identical across all three geographic samples. Put another way, the crisis has been far-reaching and is not concentrated in a small number of provinces -- the IFLS2+ sample provides a reasonably accurate portrayal of the labor market consequences of the Indonesian economic crisis. Table 1 indicates that we can safely restrict our analysis of SAKERNAS to those provinces that are also in IFLS2+ as we explore more deeply the nature and extent of the crisis. Thus, all analyses reported below use data only from the 7 provinces in IFLS2+.

Section 3: Indonesian Labor Market Before and During the Economic Crisis

What happened in urban and rural Indonesian labor markets as a result of the economic crisis? To answer that question, it is necessary to first establish the salient labor market trends in the years preceding the economic shock and then to isolate those trends that were disrupted. The best data to accomplish this goal are the annual waves of SAKERNAS spanning 1986 through 1998.

Employment rates

Table 2 presents aggregate labor market data for males and females, separating urban and rural residents. In addition to total labor force participation, the table reports employment rates in three sectors: the formal or market wage sector, self-employment, and unpaid family work. The fifth column in each panel lists mean hours worked per week, conditional on working, and the final column records mean hourly wages for those working in the market wage sector. Figure 2 graphically displays the data on market sector employment (in Panel A) and wages (in Panel B).

While rhetoric on the Indonesian labor market has alleged that the crisis has resulted in massive decreases in employment, or increases in unemployment, the data suggest otherwise. Between the mid-1980s and mid-1990s, male employment rates grew modestly in the urban sector and were relatively stable in the rural sector; there was a slight reduction in employment of males between 1997 and 1998 -- about two percentage points in urban areas and half that much in rural places. Female employment rates grew by around 25% in the urban sector during the decade prior to the crisis. That trend was reversed with the crisis: between 1997 and 1998, female employment rates declined slightly. In the rural sector, the 1980s and 1990s saw a secular decline in employment. That trend was also reversed with the crisis: the female employment rate expanded by almost three percentage points between 1997 and 1998. This suggests that one way rural families may have coped with this economic crisis is by additional work by women.

The relatively small change in total employment rates of men observed during the crisis does not imply that the disruptions in Indonesian labor markets were small. They certainly were not. Across sectors, male employment dropped by about 3.5 percentage points in the wage sector while increasing in self-employed and family jobs. This loss of male jobs in the wage sector was true in both urban and rural areas. However, it would be an over-statement to attribute the decline entirely to the crisis since market sector work peaked for rural and urban men in August 1995, well before the crisis began. The critical shift in the labor market that occurred contemporaneously with the crisis lies in the industrial distribution of jobs. In the three years prior to the crisis, there had been significant expansions in employment in manufacturing, construction, retail and sales; all of these sectors saw a dramatic reversal of that trend.⁷ Government employment remained relatively stable through the crisis. The absorbing sector was agriculture where employment expanded by 7.2 percentage points (or 20%) between August 1997 and August 1998 -- a reversal of a secular decline in the sector over the previous decade. This expansion in agricultural work took place in both rural and urban areas.

Employment in the wage sector did not contract for women living in rural areas and did so only slightly for female urban residents. The additional entry of women into the rural labor force is about

⁷Retail, construction and manufacturing grew by 25%, 22% and 11%, respectively, between 1995 and 1997; in construction and manufacturing, all of the gains were lost by 1998; half the gain was lost in the retail sector.

equally distributed between self-employed and unpaid family jobs. Like men, employment of women grew in the agriculture sector -- by about the same amount as men -- and the principal contracting sectors for women were textiles, retail, sales and services.

Column 5 of each panel reports average hours of work among those working. It has remained reasonably stable from 1986 through 1997 for all workers. Between 1997 and 1998, however, there was a substantial cut in hours worked -- about an hour per week in the urban sector and nearly two hours per week in the rural sector.

Wages

Overall, the crisis has had a small impact on aggregate employment. Wages present a sharp contrast: the drama of the crisis is reflected in the collapse of real wages of both men and women. This is vividly displayed in panel B of Figure 2 which places the real wage cuts between 1997 and 1998 in the context of a substantial rise in real wages during the previous decade. Figure 2 highlights the fact that this wage growth was especially fast in the three years preceding the crisis, particularly in urban areas. Since hours of work fell with the crisis, on average, individual earnings also collapsed during the crisis.

Trends in average wages can mask important changes in the structure of the wage distribution. It has been argued that the crisis in Indonesia has mostly affected the urban elites (Poppele et al, 1999) which would be consistent with a large decline in the mean wage but relatively modest changes in real wages of the poorest. Figure 3 summarizes the entire distribution of wages and how it has changed over time. Non-parametric estimates of the real wage distribution in the formal wage sector for 1986, 1997, and 1998 are presented. The figure demonstrates that during 1998 the *entire* distribution of male and female wages shifted sharply to the left, relative to 1997. For men, this single year shift in the wage distribution virtually wiped out all wage gains made since 1986. Among women, the leftward shift in the distribution is large but they have maintained some of their wage gains since 1986. The key point is that the crisis has affected the entire wage distribution: between 1997 and 1998, the lowest wage-earners and the highest wage-earners have all seen real wage cuts, on the order of around 40%.

Sub-group differences

To isolate those sub-groups that have been affected most by the crisis and to assess the statistical significance of the labor market changes, we have estimated a series of OLS regressions of labor market outcomes. Since employment and wage trends may be affected by shifting characteristics of the labor force, the regressions control province of residence, age (specified as splines with knots at 25, 40 and 65 years), education (stratified into five schooling categories), and urban-rural residence. Thus, the estimated year effects are standardized for changes in the attributes of workers which is potentially important given the changes in employment rates of women over the last 15 years and the substantial upgrading of the education of the labor force during this time. The models are estimated separately for men (in the left panel) and women (in the right panel). Each row of the panel represents a different sample. The first row

includes all respondents; we then stratify by sector of residence, by education of the worker and finally by age. In each stratified model, variables differentiating members of the group are obviously excluded.

Our focus is on changes between 1997 and 1998. The regressions include a calendar year dummy for every year of the survey with 1997 being omitted. Table 3 reports the coefficient on the 1998 calendar year control which measures the average (worker attribute standardized) change between 1998 and 1997 and thus highlights the impact of the economic crisis. Variance-covariance matrices are estimated by the method of the jackknife. Since SAKERNAS contains a large number of observations in each year, conventional rules of thumb for judging statistical significance are likely to be too lax. We adopt a Bayesian approach and, following Schwarz (1969), choose a critical value so that the a posteori most likely model is selected; in the linear regression case, this amounts to choosing a critical value of the test statistic that is equal to the square root of the logarithm of the number of degrees of freedom multiplied by the number of restrictions. For a t-test on a particular coefficient, in the first row of the table, the critical value is 3.66 (for employment rates) and 3.5 (for wages). We adopt this more conservative approach for judging significance in favor of a classical approach.

Several conclusions about employment can be drawn from the table. First, changes in total employment rates were relatively modest. Among men, controlling observed characteristics, there was a 1.5 percentage point decline in total labor force participation -- a decline that is concentrated in urban areas, among more educated workers and among younger workers. There is no significant impact on employment rates overall for women although, after the crisis, rural women, those with elementary schooling and older women were more likely to be working while urban, better educated women were less likely to be working.

There was more action in the sectoral distribution of work. Male employment declined by about 3.7 percentage points in the wage sector. This employment reduction was partially absorbed by the total labor force withdrawal mentioned above and partly by an increase in self-employed work and, to a lesser extent, work on the family farm. The reduction in the probability of working in the wage sector was particularly pronounced among younger men; older male workers seemed to have been better able to find self-employed jobs which presumably served to cushion the impact of the crisis.

The increases in employment rates among women discussed above are reflected in increases in self-employment and in unpaid, family work. This is true for rural women, women with elementary schooling and older women. Running counter to overall trends, employment rates of better educated women declined. For example, female employment in the highest schooling group fell by 4.2 percentage points -- an employment decline that is concentrated in the market wage sector.

For both men and women, wages fell by almost 40% between 1997 and 1998. Among male workers, no one has been spared: the wage declines are slightly higher in urban areas, uniformly distributed across education groups and are largest for the youngest. There are larger, and significant, differences in

the impact on sub-groups of women. The crisis has affected the wages of urban wage workers more, women in the middle of the education distribution have taken the biggest hit as have younger women. In fact, the only group that has not taken a very large real wage cut are women with wage-earning jobs who are age 65 to 75 -- a very select (and small) group. In sum, given the magnitude of these wage changes and the relatively small employment effects, the evidence suggests that short-run labor supply functions are fairly inelastic in Indonesia.

The statistical models reported in Table 3 provide estimates of only average effects. How were these real wage declines distributed across lower, middle, and higher wage workers? Table 4 lists the percent real wage declines that took place in the year of the crisis at selected percentiles of the wage distribution in the wage sector. To place these wage decreases in the context of the substantial rise in real wages accompanying economic growth of the previous decade, we also list real wage growth between 1986 and 1997 for the same percentiles. The distributions are arrayed separately by gender and by sector of residence.

In many ways, the Indonesian economic crisis was an equal opportunity destroyer. Among men, the substantial fall in real wages during 1998 was only slightly greater among higher wage workers than lower wage workers. This uniformity contrasts with the more beneficial effects of economic growth on male wages at the bottom of the wage distribution between 1986 and 1997. As a consequence, the economic crisis essentially wiped out the previous decade's real wage increases among higher wage men, but at least left lower wage market sector workers somewhat better off than they were in 1986. This story of uniform male wage declines during the crisis has to be amended when the data are stratified by location of employment. Very low wage male workers in urban areas had about a 7% greater wage fall than high wage urban male workers did. This contrasts with an opposite distributional pattern in rural settings where there was a tendency for higher wage rural workers to experience slightly higher wage declines. In urban and especially rural areas, the prior distributional pattern of economic wage growth differentially favored lower wage workers.

Among female wage earners, the situation was somewhat different as the hardest hit during the crisis were neither very low nor very high wage earners. Rather, those women who suffered the most were in the middle of the wage distribution. They experienced wage declines about 10% larger than those at the bottom or top of the wage distribution. As the accompanying data for 1986-1997 indicates, wages in the center of the distribution grew the most. The concentration of female wage losses in the middle of the wage distribution is a consequence of quite different patterns within the urban and rural sectors. All urban female workers at or below the median experienced about a 10% greater wage loss than urban female workers above the median. Low wage urban female workers are around median wage earners in the aggregate distribution and so it is this group that is producing the bulge in the wage shock at the center of the aggregate female wage distribution. In contrast with low wage workers taking the bigger hit in the

urban sector, among rural woman, wage earners in the upper part of the distribution experienced bigger wage declines.

A key goal of this research centers around the differential impact of the Indonesian economic crisis on workers in urban and rural labor markets. To gain some insight into this issue, it is not sufficient to simply estimate mean wage effects within each sector. Mean wages and skills are lower in rural areas so that an impact of workers of the same skill would be detected at a higher point in the rural wage distribution than in the urban wage distribution. To see whether economic shocks affect workers of similar skills differentially in rural and urban labor markets, we first arrayed (within gender) the data by percentiles in the aggregate 1997 wage distribution. We next found the corresponding percentiles in the urban and rural wage distribution that matched each aggregate percentile wage. Given that the distribution of urban wages lies above that of rural wages, these matched urban wages were found at a lower percentile and the matched rural wages at a higher percentile than the aggregate percentile wage. For each sector of residence, we then computed the percent change in the wage between 1997 and 1998 yielding a difference in rural wages and a difference in urban wages at the same real wage. The difference between the rural wage change and urban wage change is indicative of what happened during the crisis to the urban-rural wage differential among workers with the same wage in 1997. If wages are equated with skills, this "difference-in-difference" provides an indicator of the differential impact of the economic shock between urban and rural labor market at each point in the skill distribution.

Figure 4 plots the (smoothed) estimates of the urban-rural differential in the wage decline during the year of the shock separately for male and female wage earners across the 1997 wage distribution. The patterns are remarkably systematic. Among the least skilled men, there was about a 15% decrease in urban wages relative to the drop in rural wages. This greater relative wage deterioration in urban markets monotonically declines as we move up the wage (skill) distribution until there is an equal reduction of male urban and rural wages at the highest wage (skill) levels. The evidence for women essentially parallels that for men: among the lowest wage (skilled) workers, urban wages declined by between 15 to 20% more than rural wages but at the top of the wage (skill) distribution, the rural-urban wage differential did not change.

Based on this evidence, we conclude that at least in the formal wage sector, the Indonesia economic crisis hit unskilled labor markets harder in urban areas than in rural settings. In contrast, there was little geographic differentiation among the highly skilled. This difference between the change in the rural-urban wage gap between the lower skilled and higher skilled provides a "difference-in-difference-in-difference" type of test for the completeness of markets. If markets are complete (or efficient), the difference-in-difference-in-difference should be zero.⁸ The fact that it is not provides *prima facie* evidence

⁸Prices are probably lower in rural areas. However, uniformly lower prices in rural areas would only alter the location and not the shape of the curves presented in Figure 4 and so would not affect inferences based on the difference-in-difference-in-difference.

in support of the view that markets are relatively efficient among the higher skilled but that, at least in times of substantial economic turmoil, markets seem to behave as if they are incomplete among the least skilled. It may reflect barriers to mobility from the urban to rural sector among the least skilled (possibly because of liquidity constraints) or incomplete information on the part of the least skilled. In either case, the interpretation is consistent with evidence on the completeness of markets in equilibrium (Pitt and Rosenzweig, 1986; Benjamin, 1994) but incompleteness in the presence of unanticipated shocks (Murrugarra, 1998).

Section 4: The Dynamics of the Labor Market Adjustment

The previous section provided a portrait of the Indonesian labor market before and in the midst of the economic crisis based on SAKERNAS, a time series of cross-sections. While informative, these portraits are inherently limited in two key dimensions. First, SAKERNAS records income only for workers in the formal wage sector; this limitation excludes more than half of the labor force. Second, on the employment side, SAKERNAS is not informative about the dynamics of any labor market adjustments that took place as workers switched their sectoral or geographic places of work. To address these questions, we now turn to linked micro data from IFLS2 and IFLS2+. This panel survey contains information about individuals who were interviewed immediately before the onset of the 1998 Indonesian economic crisis and again a year later.

Market sector and self-employed hourly earnings

It is well-known that collection of income from self-employment in a survey setting is extremely difficult -- especially, perhaps, in a low-income and substantially agricultural setting like Indonesia. Difficulties arise, first, because of the need to calculate costs and net those out to compute profit, and, second, because incomes tend to be volatile over time, often containing an important seasonal component. The panel feature of IFLS provides some assistance. To the extent that the difficulties in measurement for a particular individual do not change between 1997 and 1998, these concerns will be somewhat mitigated; inferences based on changes in self-employment incomes over the period may not be as seriously contaminated as inferences about levels in incomes.

Figure 5 exploits information on hourly earnings in the market wage and self-employed sector reported in both IFLS2 and IFLS2+. Non-parametric estimates of the distribution of hourly earnings are reported for men and women in each sector of employment separately for 1997 and 1998.

Among men and women working in the market wage sector, the IFLS data confirm the general finding in SAKERNAS that there was a dramatic shift to the left in the distribution of real wages between 1997 and 1998. IFLS indicates there was a similar deterioration in women's income in the self-employed sector. However, in sharp contrast, real median hourly earnings of self-employed men did not change at all

during the economic crisis. Clearly, inferences about the impact of the crisis based exclusively on market sector wages (as reported in SAKERNAS) misses part of the picture.

The robustness of self-employment hourly earnings may reflect an increase in the amount of capital used in self-employment businesses (and failure of the respondents to incorporate the costs of capital in their estimates of profits). This seems unlikely given that credit markets were extremely tight at this time. It may also reflect increased labor supply of other family members when self-employment income is derived from family businesses. This is because income from the business is attributed to one family member and the estimated "hourly" earnings do not take into account the work effort of others.⁹

The distribution of hourly earnings in self-employment is fatter tailed than the distribution of wages. In part, this reflects greater error in measurement of self-employed earnings. However, Figure 5 indicates that, relative to 1997, the distribution of hourly earnings in 1998 among self-employed men was fatter tailed. This increase in variance between the years suggests that some self-employed workers benefitted from the economic crisis in terms of hourly earnings -- presumably a reflection of increases in the prices of products they sold -- while others appear to have been all but completely wiped out. In the rural sector, more than half of self-employed men actually had positive real income changes during the crisis. Since self-employed men constitute more than half the male work force, any characterizations of the economic shock in Indonesia using only the wage sector is both incomplete and misleading.

Before comparing wage and employment changes between the SAKERNAS and IFLS, a few additional adjustments are necessary to make them directly comparable. These adjustments are required because the IFLS is a panel survey while SAKERNAS is cross-sectional. This distinction mainly affects labor force patterns for either very young individuals (who are likely to enter the labor market as they age) or very old individuals (who are likely to exit as they age). To avoid contaminating the comparisons with these effects of aging, we limit our comparisons in this section to individuals who were between the ages of 22 and 64 in 1997. Since IFLS2 and IFLS2+ weights are not yet available, a more reliable basis for a comparison across the two surveys can be obtained from regression models which include as co-variables controls for the main attributes that determine the sampling weights -- province, rural-urban residence, and age of respondent. The regression model adjusted impacts reported in Tables 5 and 6 will be less influenced by different sample weights in the two surveys.

Table 5 examines changes in wages in more detail. Drawing on all respondents in IFLS who were working in either 1997 or 1998, the logarithm of wages has been regressed on an indicator variable for 1998, controlling province of residence, age and education (all measured in 1997). The regressions are

⁹This assumes only one family member is productive in the business. We have confirmed that the results presented below are not driven by that assumption by taking the opposite extreme and treating all family workers as equally productive. In that case, hourly earnings in a family business is computed as the ratio of total earnings to total number of hours worked by all family members and the same hourly earnings rate is attributed to each family worker. See Thomas, Beegle and Frankenberg (2000).

stratified by gender, sector of residence and sector of work. For each regression, the table presents the coefficient on the 1998 indicator variable which is an estimate of the percent change in real wages between 1997 and 1998 adjusting for inflation and observed worker characteristics. Mean effects, based on OLS regressions, are reported in the first row. To examine the distribution of changes in wages, quantile regressions estimates for the 25th, 50th, and 75th percentiles are reported in the following rows; t statistics are calculated by bootstrapping the quantile regressions and jackknifing OLS. (Since the jackknife is a linear approximation to the bootstrap, the two are asymptotically equivalent in the linear regression case.)

The IFLS percentage wage declines in the market sector are larger than those summarized in the previous section using SAKERNAS. In part, this reflects differences in the timing of the surveys. SAKERNAS is conducted in August each year; most of the IFLS interviews were conducted between August and December in each year. Real wages were relatively constant in 1997. Evidence from published sources suggests that in 1998 real wages continued to decline from August until December (see, for example, Papanek and Handoko, 1999). Thus, the IFLS wage changes should be larger than those in SAKERNAS.¹⁰

Figure 5 indicated that self-employment hourly earnings of males were remarkably robust during the crisis. Table 5 reveals that this is confined to males working in the rural sector. Among urban men and all women in the IFLS, self-employed hourly earnings declined as much as market sector wages.

Among workers in the market wage sector, percentage declines in wages are slightly higher in the bottom quartile of the wage distribution. In contrast, there exists a much larger negative decline at the first quartile of the self-employed wage distribution compared to the third quartile especially among men. It is common to observe that self-employment hourly earnings are more dispersed than market sector wages; not only is this true in the IFLS but, more significantly, the difference in dispersion is greater after the crisis than before.

While the regressions reported in Table 5 estimate the degree of shift at different points in wage distributions between 1997 and 1998, they do not speak directly to the distributional impacts of the crisis since it is not describing wage changes for individual workers. Wages at the first quartile of a wage distribution could decline more than wages at the third quartile for several other reasons. For example, if relatively high wage rural workers in the wage sector moved to low wage employment in the rural self-employed sector, the rural sector wage distribution would shift downward less at the top of the rural wage distribution. A more direct test of the distributional impact of the crisis on wage rates involves re-estimating the models described in Table 5 and testing whether percent wage reductions in 1998 varied by

¹⁰If interviews were randomly allocated over time in IFLS, we could control month of interview and directly compare IFLS with SAKERNAS. The interviews are not randomly assigned. Because IFLS is a panel survey, interviewers begin in areas that tend to be places migrants move from and end in areas that are destinations for migrants. This reduces the costs of tracking movers. Therefore, we do not attempt to extract any information from the month to month differences within IFLS.

education level. For the combined market and self-employed wage distributions, we found that there was no systematic pattern across respondents across the education distribution. This suggests that based on wages alone, the impact of the crisis was fairly uniform across low and high skill workers.

Employment rates and transitions

We return to the employment side of the labor market. Table 6 presents estimates of changes in sector-specific employment rates between 1997 and 1998 for IFLS and SAKERNAS. The estimates in the first column are based on OLS regressions of the probability a respondent is working in either year controlling province of residence, age, education and an indicator variable for 1998. The estimate on the 1998 indicator variable is reported in the first and second row of each panel for SAKERNAS and IFLS, respectively. The difference between the IFLS and SAKERNAS estimates is reported in the third row. The remaining three columns are based on regressions of the probability a respondent is working in a particular sector. The estimates based on IFLS are less precise than SAKERNAS because the IFLS samples are considerably smaller.

For men, the results parallel those discussed above. There are small changes in labor force participation rates of urban males with declines in activity in the market sector offset by increases in self-employment and family work. None of the differences between SAKERNAS and IFLS is significant. Among rural males, IFLS apparently overstates the decline in participation rates relative to SAKERNAS.

There is one major difference in the results for women. Relative to SAKERNAS, IFLS estimates significantly larger increases in labor force participation rates during the crisis. The difference is concentrated in a greater increase in unpaid family work in IFLS. In both urban and rural areas, there is over a four percentage point increase in the fraction of women working as unpaid family workers during the economic crisis suggesting that one way households have adapted to the economic shock has been a significant expansion of work activity by women in family businesses.

In economies such as Indonesia where formal wage sectors co-mingle alongside more traditional economic activities that often take place within the family unit, measuring work activity in surveys is not straightforward. SAKERNAS and IFLS take a somewhat different approach to this issue. Both surveys begin with the same lead-in question on whether, during the week prior to the survey, the respondent was working.¹¹ In SAKERNAS, those who answer 'yes' to that question are then asked their type of work. Family work is one option. The same screener question is asked in IFLS and those who answer 'yes' are asked about the type of work. The difference between SAKERNAS and IFLS lies in the treatment of those who answered 'no' to the lead-in question. In SAKERNAS, they are not asked any more questions about work. In IFLS, these respondents are asked a follow-up question about whether they had worked for one

¹¹Work is defined as an activity for the sake of generating income or helping to generate income that takes up at least one continuous hour.

hour for pay and, if the response is no, then they are probed by a second follow-up question about whether the respondent had worked in a family-owned business. Those who answer 'yes' to either the lead-in or the two follow-up questions about family business are treated as working and complete the employment module. This additional probing in IFLS should lead to higher estimates of labor force participation rates especially among women (which it does). But more important, this additional IFLS probing should make it easier to detect individuals switching into the family employment sector. If the increase in family work reported in IFLS arises because of the probing, the follow-up question should have captured more positive responses in IFLS2+ relative to IFLS2. It did. In IFLS2, 1.8% of female respondents who answer the question said they worked in a family business; in IFLS2+, 4.1% of women answered affirmatively. This suggests that many of the people who have entered the family business would not ordinarily consider this activity as work in some traditional sense. Given this evidence, we are inclined to find the IFLS results on female family work more credible than those in SAKERNAS.¹²

To this point, the discussion of labor force activity has been based on cross-sectional evidence. Table 7 makes use of the repeated observations on individuals in IFLS and tabulates transition rates between 1997 and 1998 for each type of work activity, conditioning in each case on the respondent's activity in 1997. The relatively stable aggregate data on male work activity derived from the cross-section comparisons masks considerable movement among the sectors and between work and non-work during the crisis. This churning is particularly prevalent in rural areas where one-third of the men and half the women changed their sector of work between 1997 and 1998.

Among rural men, the largest net transition that took place was a flow of workers from the formal market sector to self-employed jobs. 29% of rural men in wage jobs in 1997 moved to the self-employed sector by 1998. The comparable reverse flow was considerably smaller with only 15% of those in self-employment in 1997 working in the wage sector in 1998. This net flow of male employment from the wage to the self-employed sector during 1998 is consistent with our earlier finding that, in rural areas, hourly earnings of the self-employed decreased much less than of those in wage sector jobs. Among rural men, there was also a substantial flow from family work into self-employment as more than half of those rural men who were family workers in 1997 were self-employed by 1998. Note as well that there was much less movement for men between the self-employed and wage sectors in urban areas where the wage decreases in each sector were more similar.

Among women, the most salient pattern is the significant inflows into unpaid family jobs from all other sectors. In rural areas, the fraction of women working in family jobs increased from 22 to 27%

¹²See Kornis (1987) who argues that there is evidence in SAKERNAS that self-employment and family employment do not seem to mean the same thing to all interviewers. In addition, he argues that part of the year to year variation in self-employment and family employment in SAKERNAS reflects changes in interviewer quality and training over years.

between 1997 and 1998. Table 7 indicates that these net inflows into family jobs drew women from all sectors -- those not in the labor force, in the market and in the self-employed sectors. Although this net inflow into family work also characterized urban women, rates of employment in family jobs remain much lower there than in rural areas.

We have also examined which type of workers were more mobile between employment sectors. On this issue, we found that by far the least mobile workers were high wage rural workers in the formal wage sector. This is particularly clear for men.

Why are these workers so reluctant to change the sector of their employment in spite of the large decreases in their wages and the possibility of better opportunities in other sectors? The answer appears to be that they work for the government. In rural areas, two-thirds of men in the upper quartile of the market wage distribution work for the government. The comparable fraction for women is more than three-quarters.

Family income

To this point, we have analyzed Indonesian labor markets from the perspective of the individual worker. But perhaps even more so than in developed countries, the family may represent a more appropriate unit in which to measure how an economic crisis has impacted the well-being of individuals. The family represents an important institution that may help smooth out economic shocks among its members and thereby mitigate some of the risk associated with the crisis. The results obtained thus far suggest that such adjustments may well be taking place in Indonesia. The increase in female work in family businesses may be seen as a reaction to the poorer labor market prospects of men during the crisis. If so, the impact of the crisis on family earnings may be smaller than that on earnings of individual workers.

To examine this issue, Table 8 presents estimates of changes in family earnings between 1997 and 1998.¹³ The left hand panel of Table 8 contains estimates for family earnings as conventionally measured while the right hand panel augments family earnings to include the value of own-production of food.

The data summarized in the table indicate that there were considerably different impacts of the economic crisis on family earnings than that observed for individual hourly earnings (and, therefore, total earnings since, recall, hours of work declined during the crisis). This is especially clear in the rural areas. For example, Table 5 indicates that mean wages in the formal sector fell by about 60% in rural areas. The data in Table 8 indicate that mean conventional family income fell by 37% while mean family income including home production declined by 21%. While this latter decline is definitely not trivial -- it is only a third of the estimated fall in market wages in rural areas. Evaluated at the mean, the differences between

¹³Household earnings is calculated as the sum of individual earnings among all members of a household in 1997; thus, we ignore household splits, joins and new entrants. We do this in order to focus on the household unit as it existed in 1997, independent of the current location of its members.

the impact of the crisis on individual earnings and on family earnings are considerably smaller in urban areas. In those areas, at the mean, market wages fell by about 55% while family income (including the imputed value of home production) fell by 43%.

Table 8 also provides evidence on how the shape of the income distribution has changed. Recall the change in the wage distribution is dominated by a shift in location. In contrast, both the location and scale of the family income distribution has changed as reflected in the greater decline in the lower quartile, relative to the upper quartile of the distribution.

Why are our estimates of the impact of the Indonesian crisis so different for family income than for market wages in the formal wage sector? Our analysis points to two principal factors -- the lack of any significant impact of mean incomes of the self-employed in rural areas and the insurance role of the family in mitigating the full fury of this economic shock. With regard to the latter, two mechanisms suggest themselves. First, households may exploit the economies of scale associated with combining (or not splitting up). There is evidence that this is important. Mean household size increased by about 5% (Frankenberg, Smith and Thomas, 2000). Second, household members may change their allocation of times. Table 9 investigates the issues further by summarizing estimates obtained from a descriptive model of the links between (the logarithm of) family income and the number of workers in the family.

We have estimated regressions of family income on the number of family members working in the market, self employed and family business sector in that year; each family contributes two observations and both income and the number of workers are time-varying. To highlight differences associated with the crisis, the regressions include a 1998 year effect (reported in the first column of the Table 9). The number of workers in each sector is interacted with the 1998 indicator variable and those interactive effects are reported in the odd columns of the table. Column 2, for example, indicates the magnitude of the correlation between the number of family members in the market sector in 1997 and family income in 1997. Column 3 provides the difference between this correlation and the correlation between the number of market sector workers in 1998 and income in 1998. The sum of columns 2 and 3 is the correlation between number of workers in 1998 and income in 1998. Quantile models are also reported to provide a sense of how these effects differ across the income distribution.

Three main points emerge from the table. First, not surprisingly, before the crisis, the ranking of the most important contributions to family income are market wage earners, self-employed and finally family workers (who contributed not at all on average). Second, the principal impact of the crisis was to diminish the relative importance of wage earners while simultaneously increasing the importance of the self-employed and family workers. The greater role of family workers is particularly pronounced and, at the mean, it is significant in the rural sector.

Third, these differences vary both across sectors of employment and over the distribution of income. In the urban and rural sector, between 1997 and 1998, the correlation between family income and

the number of market workers changed little at median income and below but was cut in half at the top of the income distribution. The reverse is true for self-employed workers in urban areas and the change in correlation is consistently positive across the entire income distribution in rural areas. With regard to the number of family workers, the 1998 mean effects are roughly the same in rural and urban areas (being associated with about a 20% increase in income). However, in urban areas, additional family workers make a large difference at the top of the income distribution compared to their correlation with family income at the bottom. For example, we estimate a statistically insignificant negative correlation at the 25th percentile compared to a large, positive and significant correlation at the 90th percentile. In urban areas, therefore, the mitigating impact of family workers is most pronounced towards the top of the income distribution. In rural areas, the correlation between income and the number of family workers was quite different: the mitigating impact of family workers is more pronounced in the bottom half of the income distribution.

It is important to recognize that the quantile regression estimates are informative about the marginal contributions of market, self-employed and family workers at each point in the family income distribution in 1997 and 1998. They do not speak to the question of whether these workers were more (or less) important contributors to income among lower (or higher) income families. Estimation of the same models, stratifying on level of education of the household head, provides some insights into that question and indicates two aspects in which the effect of the crisis varied between poorer and less poor households. In the urban sector, the contribution of market workers to family income in 1998 was reduced far more in households whose heads had no education. In the rural sector, in 1997, the contribution of family workers was greatest among the least educated. Between 1997 and 1998, their contribution increased most among households whose heads had no schooling and those whose heads had more than primary education. As a result, in 1998, among families with little education, family workers, the self-employed and market workers were about equal contributors to family income.

Conclusion

It has been argued that the large and abrupt economic decline in Indonesia has resulted in unemployment on a massive scale. The evidence simply does not support this contention and estimates of unemployment appear to have been grossly exaggerated. In fact, labor markets in Indonesia appear to be remarkably flexible.

The drama of the financial crisis lies not in aggregate employment but in real wages, particularly in the market wage sector. While there have been increases in nominal market sector wages since the onset of the crisis, these increases have been dwarfed by the rapid increase in prices. According to SAKERNAS, real wages have collapsed by around 40% between August 1997 and August 1998. However, estimation of the exact magnitude of the decline in real wages is fraught with difficulties, some of which revolve around

computing the appropriate rate of price increases. Inflation has spiraled during this time, there is evidence of regional heterogeneity in price increases and there have been considerable shifts in relative prices.

Mindful of these concerns, there is evidence in SAKERNAS indicating wage cuts have been greater in the urban sector, on average. It turns out that this difference is entirely concentrated at the bottom of the wage distribution where the rural-urban wage differential has declined by around 15%. At the top of the wage distribution, cuts in real wages are about the same for urban and rural workers. This is *prima facie* evidence for flexible (complete) markets among the better skilled but rigidities (incomplete) markets among those with less skills, at least during times of economic disruption.

These facts emerge in two independent series of household surveys -- SAKERNAS and IFLS. SAKERNAS, however, does not record information on income from self-employment. IFLS does. Those data indicate that real hourly earnings of the self-employed in urban areas and of females in rural areas have declined by at least as much as real market sector wages. In stark contrast, real hourly earnings of self-employed males in rural areas have remained essentially stable. Given that over half the male work force are self-employed in Indonesia, conclusions about the effects of the crisis on wages that focus only on the market wage sector will substantially over-state the magnitude of the crisis.

In fact, Indonesian families have displayed a remarkable capacity for resilience in the face of the crisis. We estimate that declines in real family incomes are about half the magnitude of the declines in individual hourly earnings. To be sure, this should not be taken to imply the crisis has not been large and serious. It has. Our results do suggest, however, that households have adopted an array of strategies in an effort to mitigate the effects of the crisis.

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Figure 1: Timing of IFLS and Indonesian exchange rate

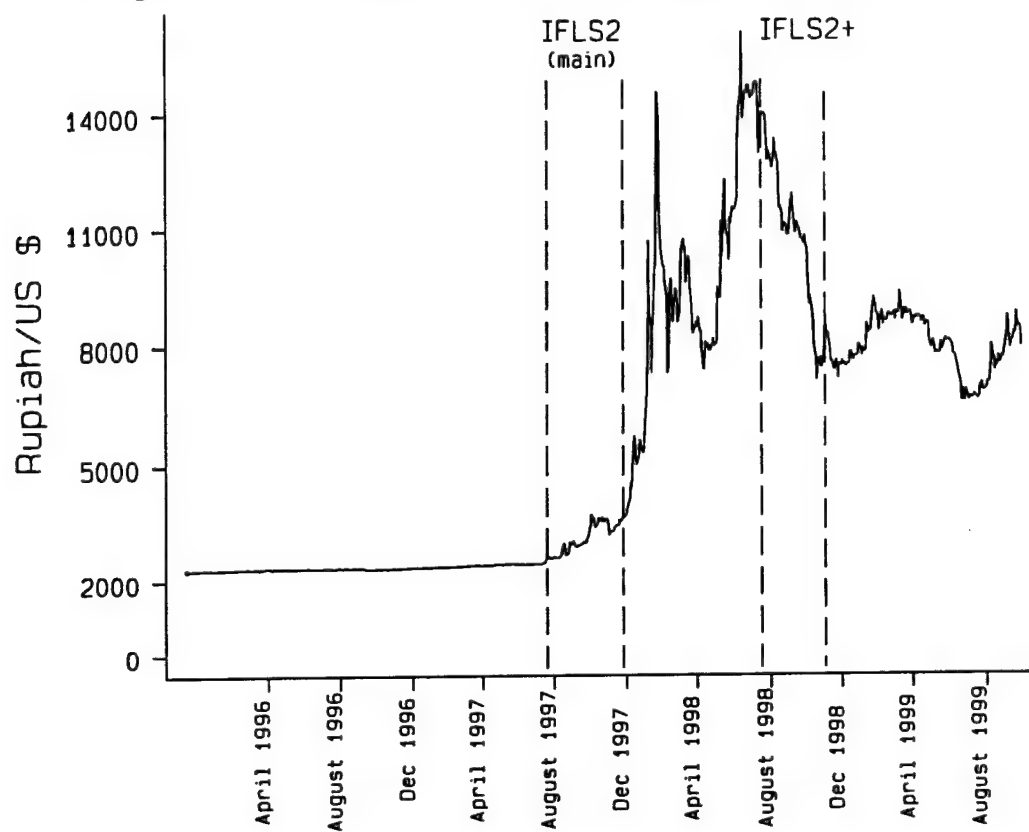


Fig 2: Employment and hourly wages by year
A: Market sector employment B: Mean hourly wages

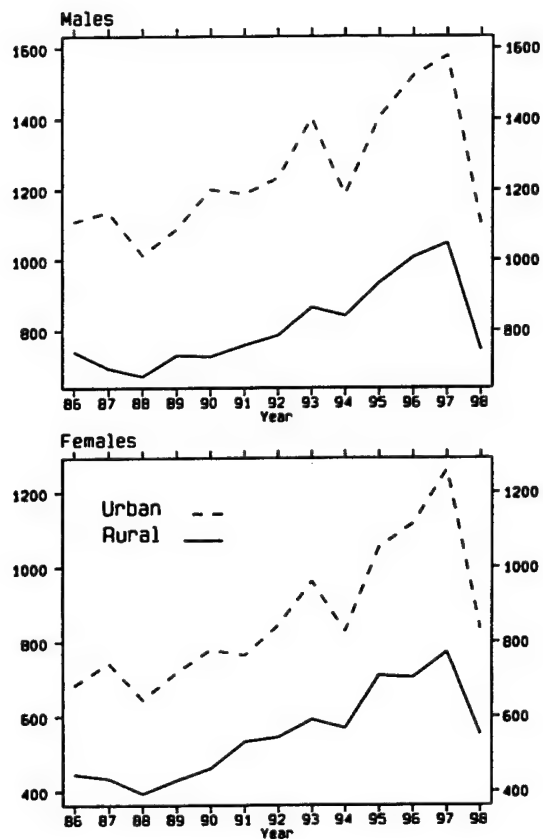
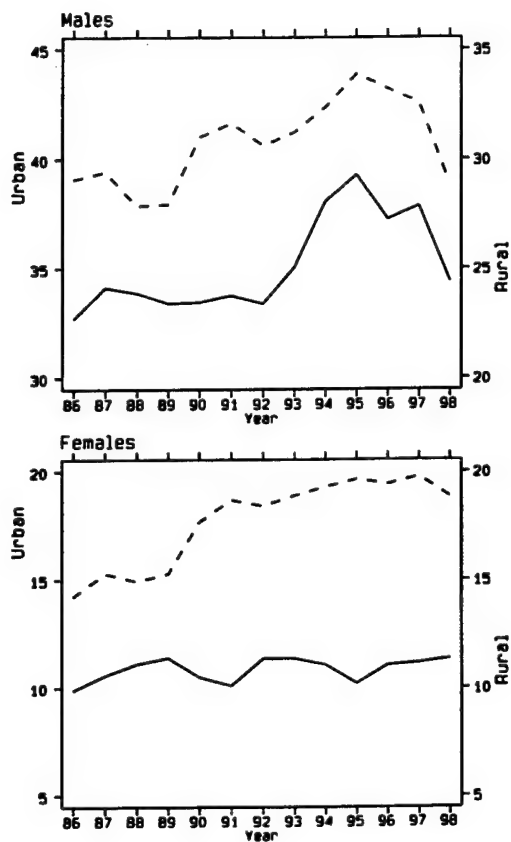
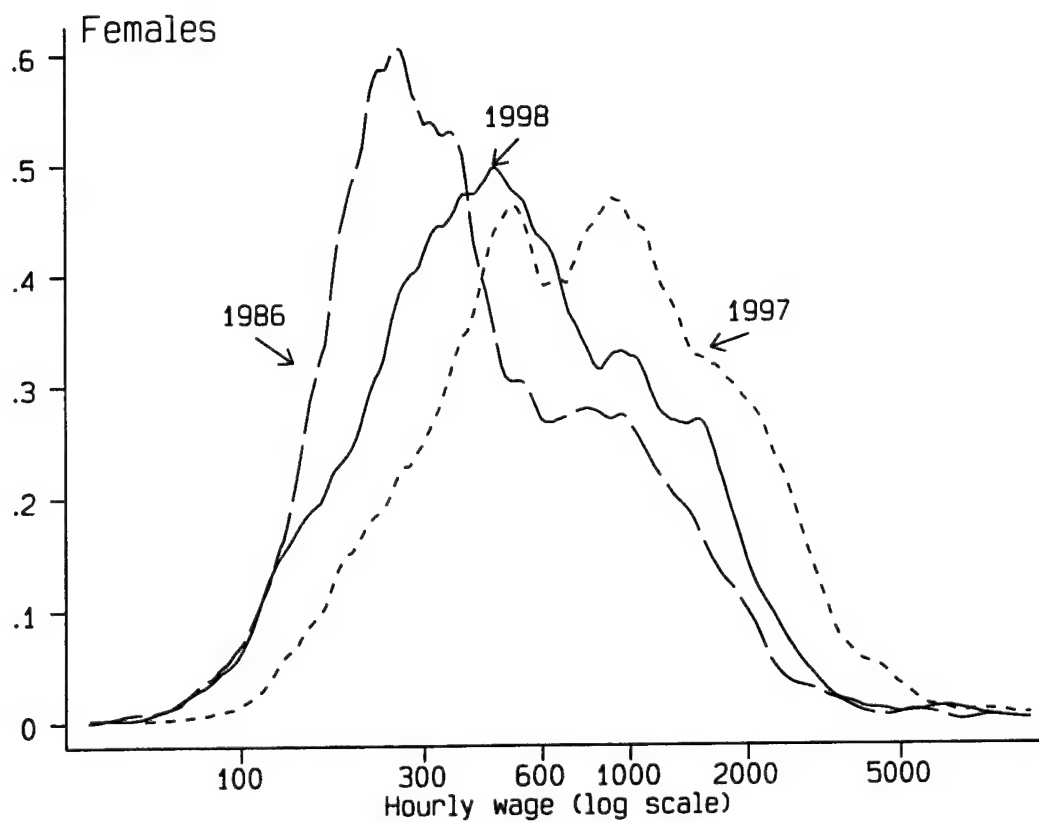
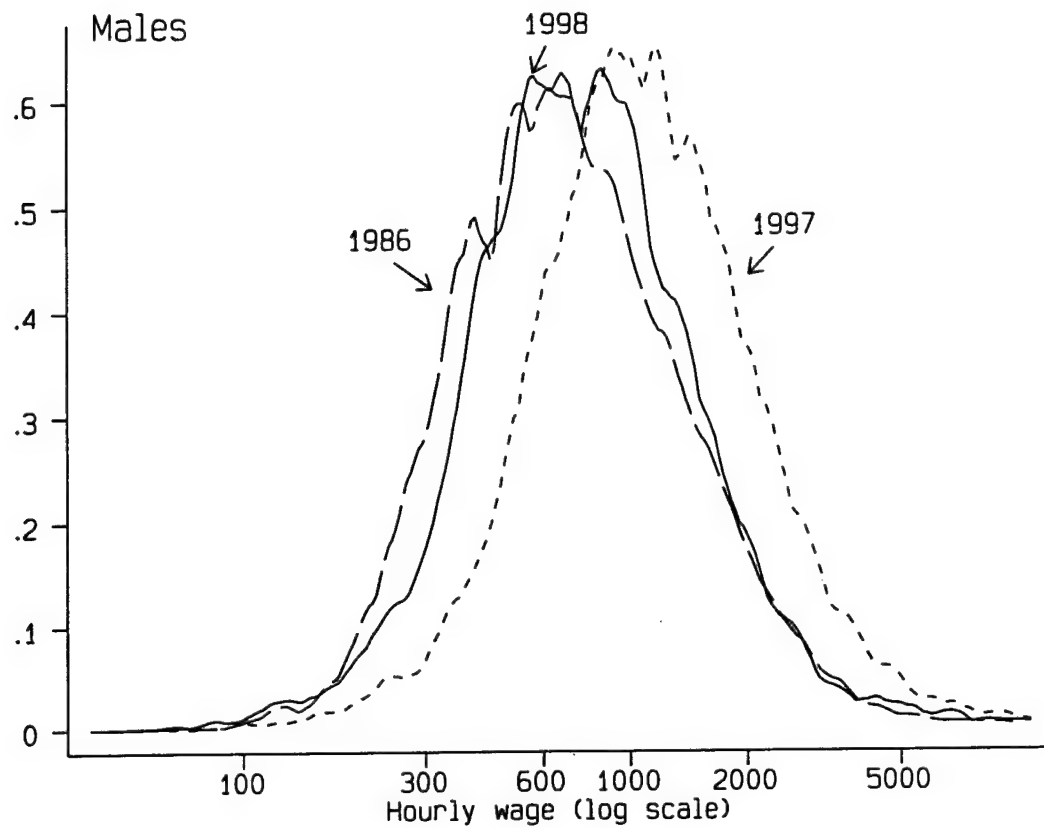


Figure 3
Hourly wage distribution: 1986, 1997, 1998
Non parametric density estimates



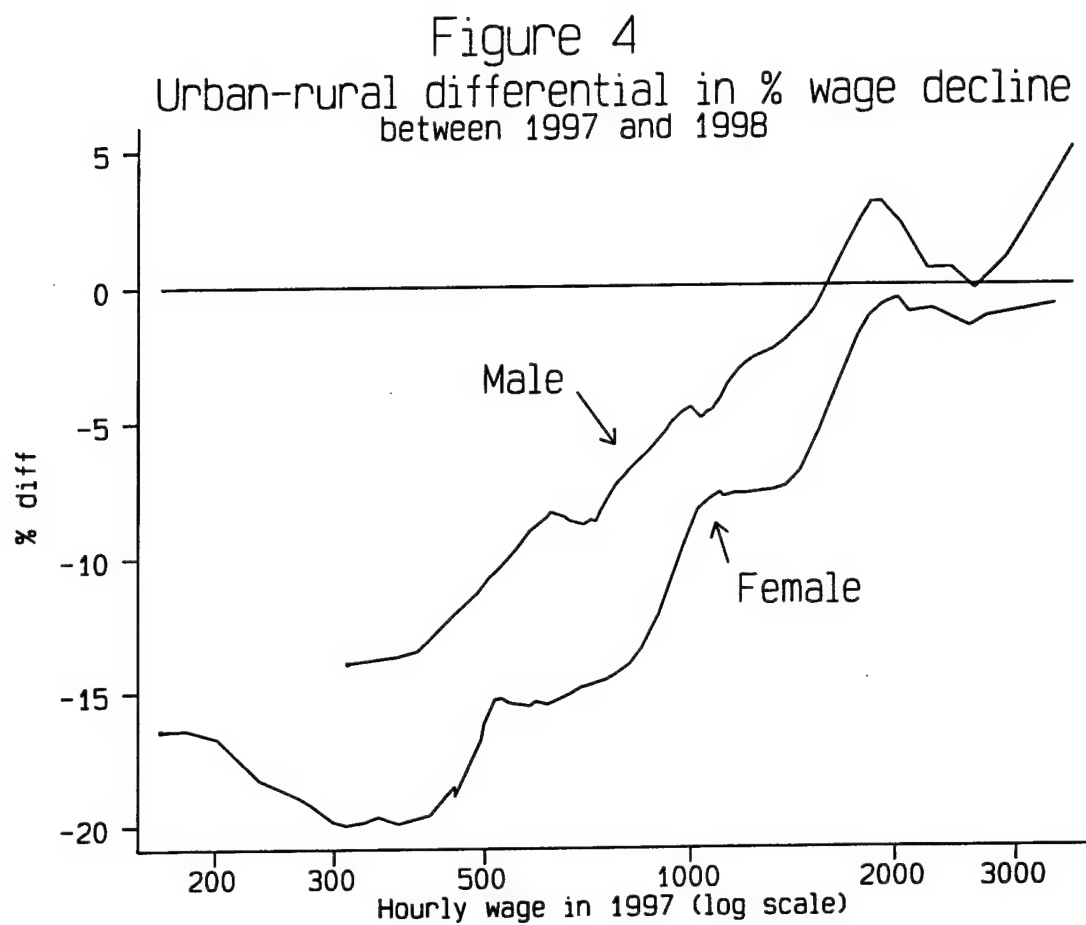
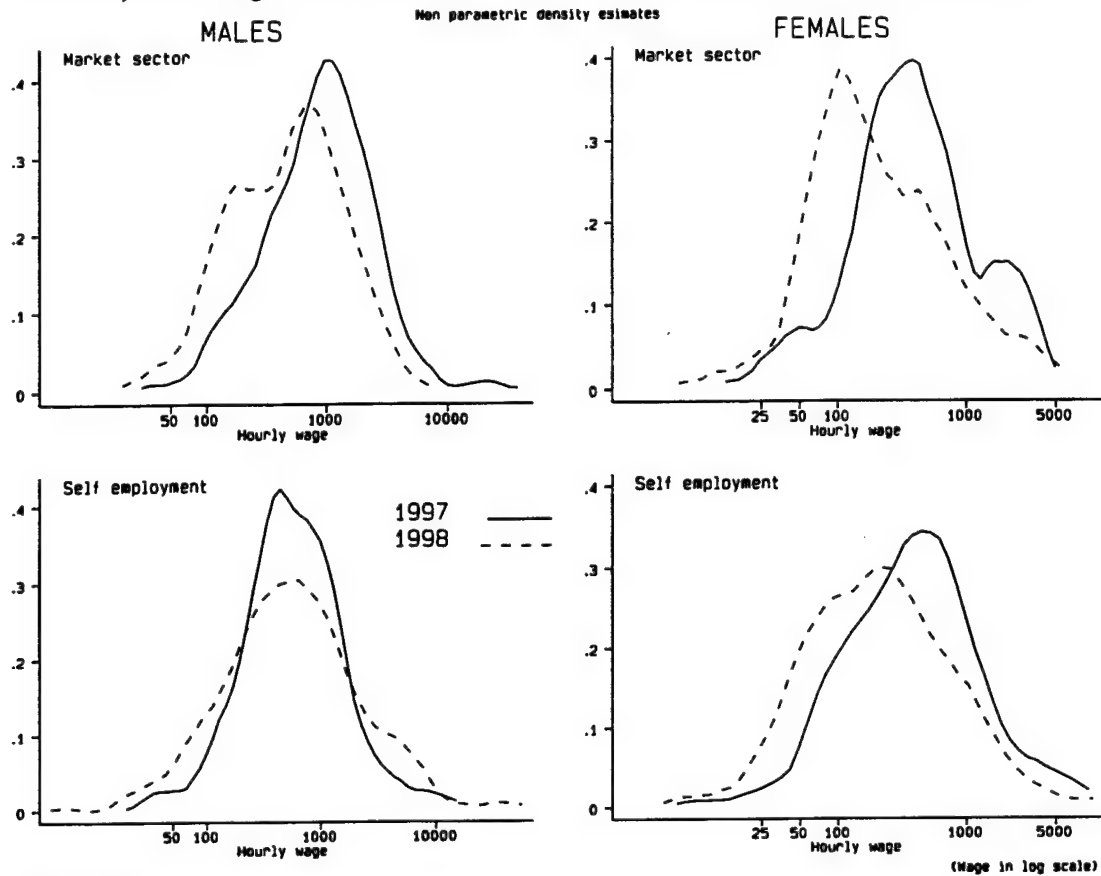


Figure 5
Hourly earnings distribution: market and self-employed sectors



Source: IFLS2/2*

Figure 6
Distribution of family income
Non parametric density estimates

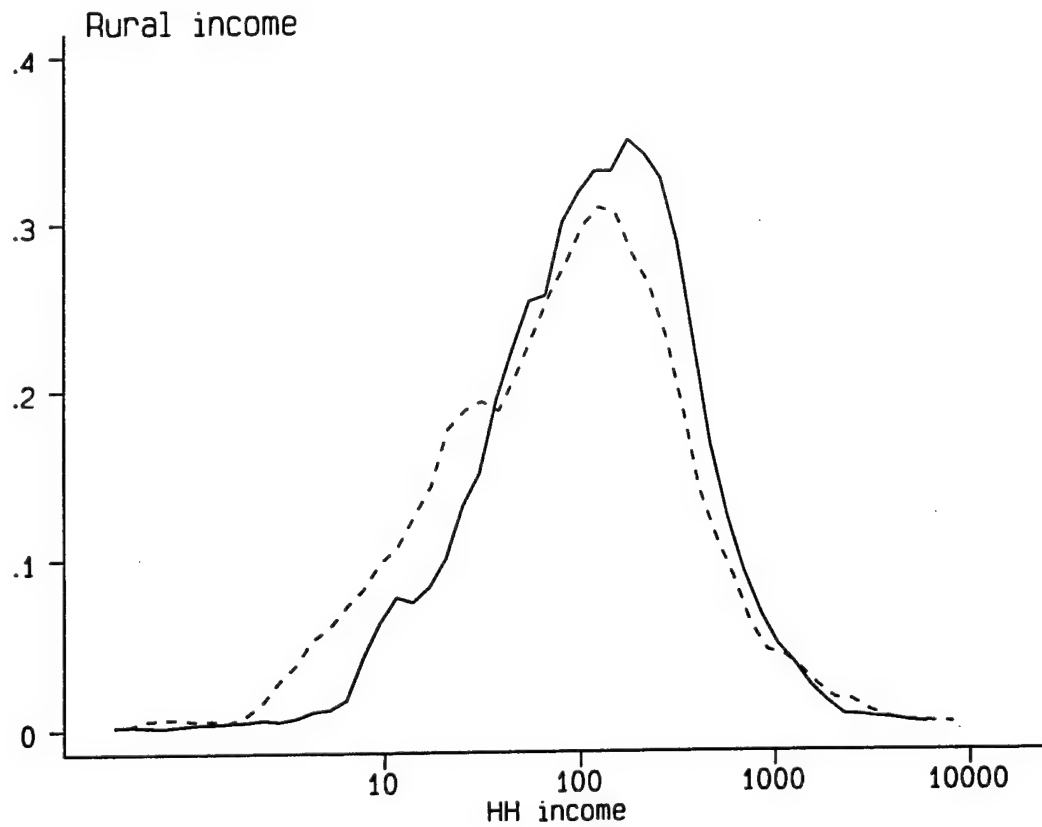
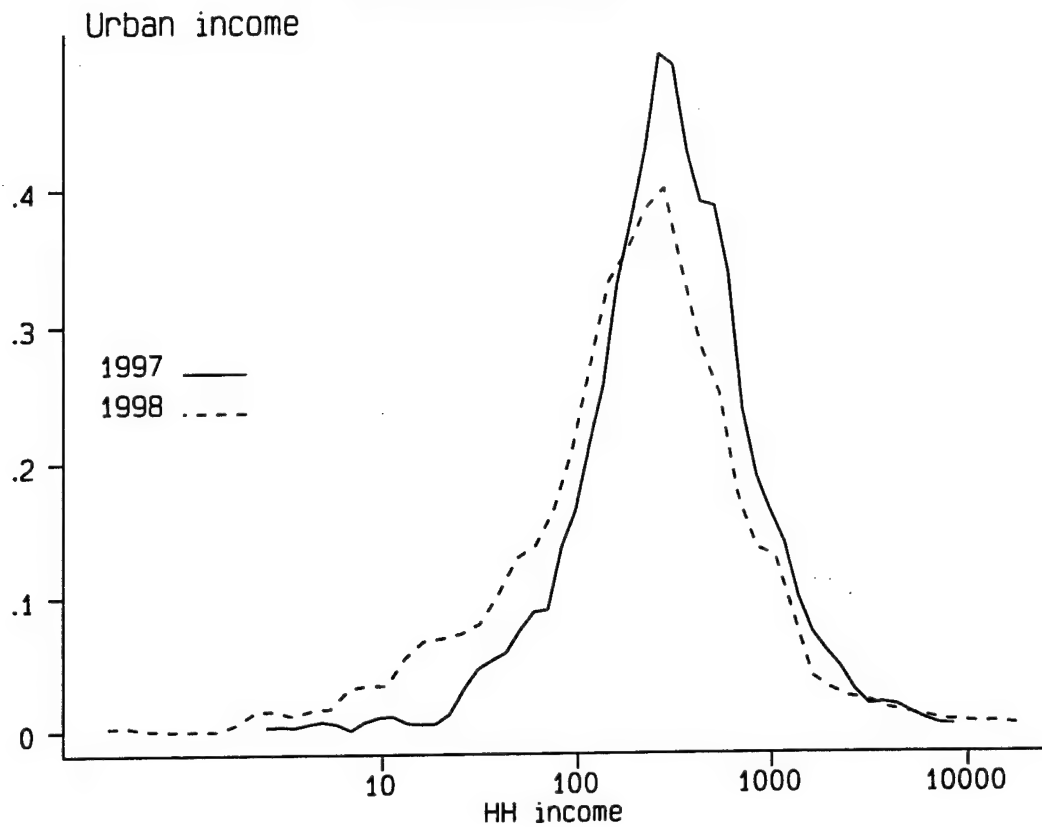


Table 1: LABOR MARKET CHANGES BETWEEN 1986, 1997 and 1998
And representativeness of IFLS provinces

| Provinces: | % change between 1986 and 1997 | | | % change between 1997 and 1998 | | |
|----------------|-----------------------------------|-------------|---------------|-----------------------------------|-------------|---------------|
| | All (1) | IFLS (2) | IFLS2+ (3) | All (1) | IFLS (2) | IFLS2+ (3) |
| MALES | | | | | | |
| Wage | 40.3 | 41.5 | 42.2 | -37.8 | -38.1 | -36.0 |
| Urban sector | 36.2 | 37.1 | 37.3 | -40.6 | -40.9 | -38.6 |
| Rural sector | 38.2 | 38.7 | 39.4 | -35.6 | -35.9 | -33.7 |
| % working | -0.5 | -0.7 | -0.3 | -0.9 | -1.1 | -1.7 |
| in wage sector | 7.1 | 7.1 | 6.4 | -3.0 | -2.7 | -2.5 |
| self-employed | -3.3 | -3.8 | -3.5 | 1.5 | 1.0 | 1.3 |
| unpaid family | -4.1 | -4.0 | -4.6 | 0.6 | 0.6 | 0.6 |
| FEMALES | | | | | | |
| Wage | 62.1 | 63.9 | 67.7 | -37.9 | -38.3 | -39.2 |
| Urban sector | 62.5 | 64.8 | 68.7 | -41.3 | -41.8 | -43.3 |
| Rural sector | 52.6 | 52.6 | 52.7 | -33.9 | -33.8 | -32.7 |
| % working | -2.3 | -0.9 | -1.3 | 1.0 | 1.3 | 1.0 |
| in wage sector | 3.7 | 3.8 | 3.8 | -0.3 | -0.2 | -0.2 |
| self-employed | 1.1 | 0.6 | 1.2 | 0.5 | 0.4 | 0.5 |
| unpaid family | -6.4 | -5.5 | -6.1 | 0.8 | 1.1 | 0.8 |

Source: SAKERNAS (1986-1998). IFLS provinces are North Sumatra, South Sumatra, West Sumatra, Lampung, Jakarta, West Java, Central Java, East Java, Yogyakarta, Bali, West Nusa Tenggara, South Kalimantan, South Sulawesi. IFLS2+ provinces are North Sumatra, South Sumatra, Jakarta, West Java, Central Java, West Nusa Tenggara and South Kalimantan. Wage changes calculated for those who earned a wage in the market wage sector only; wages measured in 1997 rupiah using BPS province-specific deflators. All estimates weighted by province weights.

Table 2: EMPLOYMENT RATES AND HOURS OF WORK

By gender, sector of residence and year

| Sector of residence /Year | MALE | | | | | FEMALE | | | | | | |
|------------------------------|-----------------|--|------------------|-------------------|-----------------------|-----------------------|-----------------|--|------------------|-------------------|-----------------------|-----------------------|
| | Any work (1) | Sector of employment Wage sector (2) | Self -emp (3) | Family wkr (4) | Hours per week (5) | Wages per hour (6) | Any work (1) | Sector of employment Wage sector (2) | Self -emp (3) | Family wkr (4) | Hours per week (5) | Wages per hour (6) |
| URBAN | | | | | | | | | | | | |
| 1986 | 66.3 | 39.1 | 24.8 | 2.5 | 46.6 | 1109 | 30.5 | 14.2 | 10.9 | 5.4 | 43.3 | 684 |
| 1987 | 67.3 | 39.4 | 24.8 | 3.1 | 46.7 | 1137 | 32.4 | 15.3 | 11.2 | 5.9 | 44.0 | 743 |
| 1988 | 65.6 | 37.8 | 24.7 | 3.0 | 46.6 | 1013 | 33.4 | 14.9 | 11.8 | 6.7 | 43.5 | 645 |
| 1989 | 65.2 | 37.9 | 24.3 | 3.0 | 47.7 | 1089 | 31.7 | 15.3 | 10.7 | 5.7 | 44.1 | 719 |
| 1990 | 69.0 | 41.0 | 24.8 | 3.2 | 47.6 | 1201 | 35.1 | 17.7 | 11.2 | 6.2 | 44.4 | 779 |
| 1991 | 70.7 | 41.6 | 26.0 | 3.1 | 48.0 | 1189 | 35.0 | 18.7 | 10.4 | 6.0 | 45.7 | 764 |
| 1992 | 69.9 | 40.6 | 26.6 | 2.7 | 47.6 | 1233 | 35.6 | 18.4 | 11.1 | 6.1 | 44.5 | 845 |
| 1993 | 70.3 | 41.2 | 26.5 | 2.6 | 47.4 | 1403 | 35.1 | 18.9 | 10.5 | 5.7 | 44.2 | 962 |
| 1994 | 71.2 | 42.4 | 26.3 | 2.6 | 46.8 | 1185 | 36.5 | 19.3 | 11.4 | 5.9 | 43.8 | 829 |
| 1995 | 72.1 | 43.9 | 26.3 | 2.0 | 47.1 | 1406 | 34.3 | 19.6 | 10.9 | 3.8 | 43.7 | 1053 |
| 1996 | 71.8 | 43.2 | 26.8 | 1.8 | 46.5 | 1522 | 36.7 | 19.4 | 13.0 | 4.3 | 42.5 | 1116 |
| 1997 | 71.9 | 42.6 | 27.2 | 2.1 | 46.8 | 1578 | 37.7 | 19.8 | 12.5 | 5.4 | 42.9 | 1260 |
| 1998 | 69.7 | 38.8 | 28.3 | 2.5 | 45.8 | 1104 | 37.0 | 18.8 | 12.2 | 5.9 | 41.9 | 835 |
| RURAL | | | | | | | | | | | | |
| 1986 | 85.8 | 22.7 | 52.0 | 11.1 | 41.1 | 740 | 56.0 | 9.9 | 17.2 | 29.0 | 31.7 | 446 |
| 1987 | 85.9 | 24.1 | 50.0 | 11.8 | 41.0 | 691 | 57.2 | 10.6 | 16.6 | 30.0 | 31.5 | 434 |
| 1988 | 85.4 | 23.9 | 49.2 | 12.3 | 41.4 | 671 | 58.4 | 11.1 | 16.6 | 30.7 | 32.2 | 394 |
| 1989 | 84.3 | 23.4 | 48.7 | 12.2 | 41.8 | 730 | 58.6 | 11.4 | 17.2 | 30.0 | 33.1 | 429 |
| 1990 | 86.5 | 23.4 | 51.2 | 11.8 | 41.5 | 727 | 55.5 | 10.5 | 16.4 | 28.6 | 32.1 | 462 |
| 1991 | 86.2 | 23.7 | 51.7 | 10.8 | 41.5 | 759 | 54.5 | 10.1 | 16.2 | 28.1 | 32.1 | 533 |
| 1992 | 86.4 | 23.4 | 52.0 | 11.0 | 40.4 | 786 | 56.6 | 11.3 | 16.1 | 29.2 | 31.4 | 545 |
| 1993 | 85.7 | 25.1 | 51.2 | 9.5 | 40.8 | 866 | 54.4 | 11.3 | 16.7 | 26.3 | 32.2 | 592 |
| 1994 | 85.3 | 28.1 | 48.9 | 8.3 | 40.8 | 842 | 53.6 | 11.1 | 18.9 | 23.6 | 33.0 | 568 |
| 1995 | 84.7 | 29.3 | 47.5 | 7.9 | 41.4 | 935 | 45.2 | 10.2 | 17.2 | 17.8 | 32.5 | 710 |
| 1996 | 85.4 | 27.2 | 52.2 | 5.9 | 39.5 | 1007 | 53.0 | 11.0 | 23.8 | 18.1 | 30.7 | 705 |
| 1997 | 85.0 | 27.9 | 49.8 | 7.3 | 40.7 | 1048 | 51.7 | 11.2 | 18.6 | 21.9 | 32.0 | 772 |
| 1998 | 84.0 | 24.4 | 51.5 | 8.1 | 38.6 | 745 | 54.6 | 11.3 | 19.9 | 23.4 | 30.5 | 552 |

Source: SAKERNAS using IFLS2+ provinces only. Hourly wages in 1997 Rp deflated with BPS province-specific price indices. All rates and means weighted.

Table 3: MAGNITUDE OF EMPLOYMENT AND WAGE SHOCKS
Regression estimates of difference in year effects between 1997 and 1998
By gender, sector of residence, level of education and age.

| | MALE | | | | | FEMALE | | | | |
|---------------------------|----------------------|----------------|-----------------|------------------|------------------------|----------------------|-----------------|-----------------|-----------------|------------------------|
| | Sector of employment | | Wages | | Wages per hr (5) | Sector of employment | | Wages | | Wages per hr (5) |
| | Any (1) | Self (3) | Family (4) | Wage (2) | | Any (1) | Self (3) | Family (4) | Wage (2) | |
| All country | -1.47 (5.70) | 1.74 (5.31) | 0.47 (2.81) | -3.68 (10.87) | -38.46 (53.17) | 0.81 (2.24) | 0.50 (1.84) | 0.89 (3.40) | -0.59 (2.26) | -38.99 (33.42) |
| Sector | | | | | | | | | | |
| Urban | -2.74 (6.94) | 1.08 (2.39) | 0.39 (2.25) | -4.21 (8.39) | -39.99 (43.71) | -1.16 (2.26) | -0.35 (0.99) | 0.51 (1.99) | -1.32 (3.22) | -43.28 (30.28) |
| Rural | -0.34 (1.04) | 2.41 (5.15) | 0.52 (1.88) | -3.27 (7.24) | -36.08 (30.98) | 2.51 (5.01) | 1.23 (2.98) | 1.20 (2.70) | 0.07 (0.22) | -31.20 (15.96) |
| Education | | | | | | | | | | |
| None | -0.92 (0.72) | 1.61 (0.97) | 0.52 (0.91) | -3.05 (2.38) | -38.23 (7.19) | 1.40 (1.26) | 0.20 (0.20) | -0.02 (0.02) | 1.21 (1.74) | -30.38 (6.17) |
| Elementary incomplete | 1.32 (2.31) | 2.77 (3.05) | 0.19 (0.51) | -1.64 (1.97) | -33.06 (14.63) | 4.36 (5.18) | 1.42 (1.91) | 1.98 (2.81) | 0.96 (1.77) | -25.12 (7.37) |
| Elementary complete | -0.68 (1.79) | 2.73 (4.68) | 0.61 (1.94) | -4.02 (7.06) | -37.56 (28.94) | 2.55 (4.15) | 1.17 (2.45) | 0.81 (1.69) | 0.57 (1.39) | -40.51 (18.09) |
| Jr Secondary | -2.26 (3.76) | 0.50 (0.73) | 0.43 (1.02) | -3.20 (4.42) | -39.87 (23.17) | -1.19 (1.47) | -0.03 (0.06) | 0.36 (0.65) | -1.51 (2.72) | -55.37 (16.64) |
| Sr Secondary/ Tertiary | -3.40 (6.13) | 0.72 (1.22) | 0.49 (1.76) | -4.60 (6.55) | -39.44 (33.18) | -4.15 (4.83) | -1.24 (2.64) | 1.18 (2.88) | -4.09 (4.97) | -39.46 (23.40) |
| Age | | | | | | | | | | |
| 15-24 | -2.08 (3.05) | 0.94 (2.04) | 1.00 (2.12) | -4.02 (6.80) | -38.50 (23.31) | -0.86 (1.33) | 0.37 (1.20) | 0.79 (1.83) | -2.02 (3.93) | -44.16 (22.14) |
| 25-39 | -1.89 (5.57) | 2.26 (3.68) | 0.52 (2.26) | -4.67 (7.43) | -41.07 (40.40) | 0.66 (1.07) | 0.02 (0.05) | 0.75 (1.67) | -0.11 (0.24) | -39.68 (22.69) |
| 40-64 | -1.24 (3.71) | 1.61 (2.41) | -0.11 (0.79) | -2.74 (4.19) | -33.75 (25.40) | 2.27 (3.21) | 1.32 (2.00) | 1.05 (1.83) | -0.10 (0.22) | -30.71 (12.12) |
| 65-75 | 1.74 (1.50) | 3.75 (3.05) | -0.28 (0.79) | -1.73 (2.15) | -31.57 (5.89) | 3.66 (2.98) | 0.97 (0.90) | 1.24 (1.55) | 1.45 (2.63) | -13.29 (1.55) |

Source: SAKERNAS. Based on OLS regressions including individual respondents in all years 1986 through 1998 resident in IFLS2+ provinces. Regressions control province of residence, sector of residence, age and education of respondent. Percentage point change between 1997 and 1998 reported for employment rates; percentage changes reported for wages. Associated t statistic reported in parentheses below each estimated change. Variance-covariance matrix estimated by method of jackknife. Estimates for all country (in first row) based on 638,836 observations (male employment), 194,379 observations (male wages) 658,259 observations (female employment), 85,573 observations (female wages). Critical value of t statistic for significance based on Schwarz criterion is 3.66 for employment rates and 3.5 for wage functions in these cases.

Table 4: DISTRIBUTION OF WAGE SHOCKS

Short-term and longer-term percent wage changes across percentiles of wage distribution

| Wage Pctile | Urban & Rural | | Urban | | Rural | |
|----------------|---------------|--------------|--------------|--------------|--------------|--------------|
| | % change btwn | | %change btwn | | %change btwn | |
| | 86&97 (1) | 97&98 (2) | 86&97 (1) | 97&98 (2) | 86&97 (1) | 97&98 (2) |
| MALES | | | | | | |
| 10 | 47.8 | -35.9 | 47.0 | -41.6 | 41.4 | -31.6 |
| 30 | 49.0 | -35.9 | 37.9 | -39.9 | 43.0 | -33.0 |
| 50 | 43.6 | -36.8 | 31.1 | -37.3 | 43.0 | -34.3 |
| 70 | 36.9 | -36.7 | 31.0 | -36.4 | 35.6 | -32.9 |
| 90 | 33.5 | -36.4 | 34.4 | -36.3 | 27.1 | -40.1 |
| 95 | 37.7 | -37.4 | 38.7 | -34.0 | 20.9 | -34.0 |
| FEMALES | | | | | | |
| 10 | 50.9 | -36.0 | 68.5 | -47.8 | 34.4 | -28.1 |
| 30 | 69.3 | -40.6 | 81.0 | -47.1 | 49.9 | -30.3 |
| 50 | 81.0 | -46.5 | 71.7 | -50.6 | 55.1 | -31.3 |
| 70 | 77.0 | -42.2 | 50.7 | -42.3 | 58.7 | -34.4 |
| 90 | 61.6 | -38.3 | 51.2 | -38.1 | 59.7 | -37.7 |
| 95 | 53.6 | -38.2 | 49.0 | -37.2 | 47.4 | -35.1 |

Source: SAKERNAS for IFLS2+ provinces. Estimates in column (1) are $\ln(\text{wage})$ in 1997 evaluated at appropriate percentile of 1997 wage distribution - $\ln(\text{wage})$ in 1986 evaluated at same percentile of 1986 wage distribution. Estimate in column (2) is $\ln(\text{wage})$ in 1998 - $\ln(\text{wage})$ in 1997 evaluated at appropriate percentiles.

Table 5: ADJUSTED HOURLY EARNINGS SHOCKS IN MARKET AND SELF-EMPLOYED SECTORS
 Estimated percentage change in earnings per hour between 1997 and 1998
 IFLS panel respondents

| Quantile | Male | | | | Female | | | |
|----------|------------------------|----------------------|------------------------|----------------------|------------------------|----------------------|------------------------|----------------------|
| | Urban Market (1) | Urban Self (2) | Rural Market (3) | Rural Self (4) | Urban Market (5) | Urban Self (6) | Rural Market (7) | Rural Self (8) |
| Mean | -53.5 [9.5] | -66.9 [5.3] | -57.1 [6.8] | -11.2 [1.2] | -58.6 [6.2] | -51.2 [3.8] | -58.6 [4.5] | -55.8 [4.8] |
| 25%ile | -57.5 [6.4] | -89.0 [5.2] | -59.7 [5.6] | -24.9 [1.8] | -66.5 [4.4] | -76.7 [3.6] | -76.1 [4.5] | -63.3 [3.3] |
| 50%ile | -46.0 [6.7] | -57.7 [3.4] | -49.8 [3.6] | -10.6 [1.1] | -55.9 [5.2] | -56.0 [3.7] | -65.7 [4.4] | -48.8 [3.1] |
| 75%ile | -40.3 [6.4] | -45.2 [2.5] | -46.6 [5.2] | 1.2 [0.1] | -54.0 [5.8] | -60.8 [3.6] | -48.1 [2.9] | -53.6 [3.6] |
| # obs | 773 | 390 | 486 | 611 | 326 | 358 | 210 | 458 |

Source: IFLS2 and IFLS2+ Panel respondents. Quantile regression and OLS regression estimates based on panel respondents interviewed in IFLS2 (1997) and IFLS2+ (1998). Estimated percentage difference in hourly earnings in 1998 and hourly earnings in 1997 reported in table; [t statistics] in parentheses estimated by method of jackknife (for mean) and bootstrap (for quantiles). Regressions control age, education and province of residence of respondent.

Table 6: EMPLOYMENT SHOCKS
Adjusted Differences Between 1997 and 1998
SAKERNAS and IFLS

| | All Workers (1) | Market Sector (2) | Self- employed (3) | Family Workers (4) |
|-------------------------------|-----------------------|-------------------------|--------------------------|--------------------------|
| MALES | | | | |
| <i>URBAN</i> | | | | |
| SAKERNAS | -2.9 [6.4] | -4.3 [6.9] | 1.2 [2.1] | 0.2 [1.3] |
| IFLS | -1.2 [0.7] | -3.9 [1.6] | 1.5 [0.6] | 1.2 [2.0] |
| Difference (IFLS-SAKERNAS) | 1.7 [1.0] | 0.4 [0.2] | 0.3 [0.1] | 1.0 [1.6] |
| <i>RURAL</i> | | | | |
| SAKERNAS | -0.6 [2.1] | -3.3 [6.1] | 2.3 [4.1] | 0.4 [1.7] |
| IFLS | -3.4 [3.0] | -0.6 [0.3] | -1.6 [0.7] | -1.3 [1.4] |
| Difference | -2.8 [2.4] | 2.8 [1.3] | -3.9 [1.7] | 1.7 [1.8] |
| FEMALES | | | | |
| <i>URBAN</i> | | | | |
| SAKERNAS | -0.5 [0.8] | -0.5 [1.0] | -0.5 [1.0] | 0.5 [1.5] |
| IFLS | 6.6 [2.8] | 1.0 [0.5] | 1.5 [0.7] | 4.2 [3.9] |
| Difference | 7.1 [2.9] | 1.4 [0.7] | 1.9 [0.9] | 3.7 [3.3] |
| <i>RURAL</i> | | | | |
| SAKERNAS | 3.1 [5.4] | 0.3 [0.8] | 1.5 [3.0] | 1.3 [2.4] |
| IFLS | 7.0 [3.4] | 1.1 [0.8] | 0.3 [0.1] | 5.7 [3.1] |
| Difference | 3.9 [1.8] | 0.8 [0.5] | -1.2 [0.7] | 4.4 [2.3] |

Source: SAKERNAS and IFLS. OLS regression estimates of difference between 1997 and 1998; asymptotic t statistics in parentheses estimated by method of jackknife. Regressions include controls for age, education and province of residence.

Table 8: FAMILY INCOME SHOCKS

Estimated percentage change in family income between 1997 and 1998
IFLS panel households

| Quantile | Family earnings from market and self- employment work | | Family earnings and imputed value of home production | |
|----------|---|--------------|--|--------------|
| | Sector of residence in 1997 | | Sector of residence in 1997 | |
| | Urban | Rural | Urban | Rural |
| Mean | -47 [10.4] | -37 [8.0] | -43 [8.8] | -21 [5.0] |
| 25%ile | -55 [7.0] | -53 [8.4] | -64 [8.6] | -33 [6.0] |
| 50%ile | -34 [10.1] | -32 [4.4] | -39 [9.2] | -22 [4.2] |
| 75%ile | -23 [5.0] | -20 [4.0] | -23 [5.4] | -20 [5.4] |
| 90%ile | -17 [2.5] | -5 [0.6] | -22 [2.3] | -14 [2.1] |

Source: IFLS2/2+ panel households. Estimates of %age decline in family income between 1997 and 1998 derived from OLS and quantile regressions; [t statistics] in parentheses based on jackknife (for OLS) and bootstrap (for quantile regressions).

Table 9: FAMILY INCOME SHOCKS AND FAMILY EMPLOYMENT

Correlations between number of family members employed in each sector and family income

Main effects and interactions with control for 1998

IFLS panel households

| | # workers in each year: Yr=1998 (1) | market sector (2) | sector *1998 (3) | self employment (4) | *1998 (5) | unpaid family bus (6) | *1998 (7) |
|--------------|---|----------------------|------------------------|------------------------|---------------|--------------------------|---------------|
| <i>URBAN</i> | | | | | | | |
| Mean | -50.9 [4.0] | 44.1 [11.0] | -5.8 [0.9] | 24.4 [5.2] | 3.3 [0.4] | -6.3 [0.6] | 22.5 [1.2] |
| 25%ile | -73.3 [3.0] | 44.6 [7.5] | 1.6 [0.1] | 22.3 [3.5] | -2.7 [0.3] | 4.0 [0.3] | -8.1 [0.4] |
| 50%ile | -36.7 [2.3] | 41.5 [9.6] | 0.7 [0.1] | 20.2 [3.5] | -6.8 [0.6] | -3.4 [0.3] | 2.8 [0.2] |
| 75%ile | -19.5 [1.3] | 44.6 [10.9] | -10.3 [1.6] | 24.5 [3.6] | 3.7 [0.3] | -11.6 [0.7] | 26.1 [1.0] |
| 90%ile | -12.3 [0.5] | 36.5 [5.8] | -15.1 [1.4] | 24.7 [2.4] | 18.4 [1.2] | -18.2 [0.6] | 92.8 [2.0] |
| <i>RURAL</i> | | | | | | | |
| Mean | -53.7 [3.5] | 53.4 [8.9] | -14.1 [1.6] | 19.4 [3.2] | 10.8 [1.2] | -5.6 [0.9] | 27.6 [2.9] |
| 25%ile | -85.3 [3.9] | 58.8 [5.9] | -5.1 [0.4] | 19.7 [2.4] | 10.1 [0.7] | -6.2 [0.8] | 33.6 [2.9] |
| 50%ile | -40.8 [1.7] | 54.1 [8.8] | -18.7 [1.9] | 19.0 [2.8] | 2.0 [0.2] | -1.6 [0.2] | 21.3 [1.7] |
| 75%ile | -33.9 [2.2] | 41.6 [5.4] | -10.6 [1.1] | 11.7 [1.5] | 11.9 [1.1] | -1.9 [0.2] | 14.7 [1.4] |
| 90%ile | -8.2 [0.3] | 42.4 [3.6] | -27.4 [1.5] | 18.4 [1.9] | 9.3 [0.6] | -7.3 [1.1] | 14.2 [0.8] |

Source: IFLS2/2+ panel households. Estimates based on regressions of ln(family income) on indicator variable for 1998 (column 1), number of workers in market sector, self employment and family work in each year (even columns) plus interaction between the number of workers in each sector and 1998 indicator variable (odd columns). [t statistics] in parentheses based on jackknife for OLS and bootstrap for quantile regressions.